

Notes on Baldi chapter 1¹

1 Basic Properties of Probability

Definition. A probability measure on $(\Omega, \mathcal{F})$ is a function $P: \mathcal{F} \rightarrow \mathbb{R}_+$ which is normalized and σ -additive, that is $P(\Omega) = 1$ and $P(\cup_{n \geq 1} A_n) = \sum_{n \geq 1} P(A_n)$ for disjoint countable families in $\mathcal{F}$.

Recall that $A \setminus B \equiv A \cap B^c$. We will abbreviate $A \cap B$ as AB whenever not confusing.

Proposition. Let P be a probability measure on $(\Omega, \mathcal{F})$. Then

1. $P(\emptyset) = 0$
2. P is finitely additive: $P(\cup_{n=1}^N A_n) = \sum_{n=1}^N P(A_n)$ for any finite disjoint family in $\mathcal{F}$, and in particular $P(A^c) = 1 - P(A)$
3. $P(A) = P(AB) + P(AB^c)$ for all $A, B \in \mathcal{F}$
4. P is monotone: $P(A) \leq P(B)$ if $A \subseteq B$, and in particular $P(A) \leq 1$ for all $A \in \mathcal{F}$.
5. $P(A \cup B) = P(A) + P(B) - P(AB)$ for all $A, B \in \mathcal{F}$
6. P is continuous: if $(A_n)_{n \geq 1}$ is an increasing sequence (that is $A_n \subseteq A_{n+1}$ for all n) with union $\cup_{n \geq 1} A_n = A$, or it is a decreasing sequence ($A_{n+1} \subseteq A_n$ for all n) with intersection $\cap_{n \geq 1} A_n = A$, then

$$P(A) = \lim_{n \rightarrow \infty} P(A_n).$$

Proof. 1. Since the sequence with $A_n = \emptyset$ for all n is a disjoint family and in this case $\cup_{n \geq 1} A_n = \emptyset$, by σ -additivity we have $P(\emptyset) = \sum_{n \geq 1} P(\emptyset)$, which can be true only if $P(\emptyset) = 0$.

For 2 it suffices to take $A_n = \emptyset$ for all $n > N$. For the second assertion in 2 simply observe that $\Omega = A \cup A^c$, a disjoint union, and that $P(\Omega) = 1$.

3. $\Omega = B \cup B^c$ so $A = A\Omega = AB \cup AB^c$, disjoint union, therefore 2 implies the result.

4. If $A \subseteq B$ then $BA = A$ whence $B = A \cup BA^c$, another disjoint union; this and non-negativity of P give the result. The second assertion follows from $A \subseteq \Omega$.

5. $A \cup B = AB^c \cup AB \cup BA^c$ (disjoint), but from 3 we have $P(A) = P(AB) + P(AB^c)$ and $P(B) = P(AB) + P(BA^c)$; substituting the values of $P(AB^c)$ and $P(BA^c)$ we get the result.

6. Consider first an increasing sequence. Obviously $A_2 = A_2 A_1 \cup A_2 A_1^c = A_1 \cup (A_2 \setminus A_1)$ (disjoint) and by an easy induction argument we get $A_n = \cup_{i=1}^n (A_i \setminus A_{i-1})$, with $A_0 \equiv \emptyset$, union again disjoint. We also have $A = \cup_{i \geq 1} (A_i \setminus A_{i-1})$, disjoint; indeed obviously $\cup_{i \geq 1} (A_i \setminus A_{i-1}) \subseteq \cup_{i \geq 1} A_i$; on the other hand suppose $\omega \in A$ and let k be the smallest integer such that $\omega \in A_k$; then $\omega \in A_k \setminus A_{k-1} \subseteq \cup_{i \geq 1} (A_i \setminus A_{i-1})$. Using σ -additivity we then

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conclude that

$$\begin{aligned} P(A) &= P(\cup_{i \geq 1} (A_i \setminus A_{i-1})) = \sum_{i \geq 1} P(A_i \setminus A_{i-1}) = \lim_{n \rightarrow \infty} \sum_{i=1}^n P(A_i \setminus A_{i-1}) \\ &= \lim_{n \rightarrow \infty} P(\cup_{i=1}^n (A_i \setminus A_{i-1})) = \lim_{n \rightarrow \infty} P(A_n). \end{aligned}$$

For a decreasing sequence observe that $A^c = \cup_{n \geq 1} A_n^c$ and $(A_n^c)_{n \geq 1}$ is an increasing sequence; then from what we just showed $1 - P(A_n) = P(A_n^c) \rightarrow P(A^c) = 1 - P(A)$, which implies the last result. $\square$

The fact that $P(A) = P(AB) + P(AB^c)$ will be used all the time. The continuity of probability is central to most developments of probability theory.

Remark. Observe that when Ω is finite, $\Omega = \{\omega_1, \dots, \omega_n\}$, to define a probability P on Ω (with $\mathcal{F} = 2^\Omega$) all we need to do is specify n non-negative numbers $p_i = P(\{\omega_i\})$ with $\sum_{i=1}^n p_i = 1$. Then for any $A \subseteq \Omega$ we have $P(A) = \sum_{i: \omega_i \in A} p_i$.

2 Conditional Probability and Independence

The definition of conditional probability is, with $P(B) > 0$,

$$P(A|B) := \frac{P(AB)}{P(B)}.$$

The natural definition of independence is the following:

Definition. Let $A, B \in \mathcal{F}$. Event A is independent of B if $P(A|B) = P(A)$.

This says that knowing that B occurred does not influence the probability of A . If this is true, could it be that - asymmetrically - knowledge of A does influence the probability of B ? The answer is no:

Proposition. *The following are equivalent:*

- (i) A is independent of B
- (ii) B is independent of A , that is $P(B|A) = P(B)$
- (iii) $P(AB) = P(A) \cdot P(B)$

Proof. The equivalence (i) $\iff$ (iii) follows immediately from the definition, since

$$\frac{P(AB)}{P(B)} = P(A) \iff P(AB) = P(A) \cdot P(B)$$

and from this the other equivalence also follows immediately. $\square$

Thus the concept of independence of two event is really symmetrical, and we may use as alternative definition “ A and B are independent if $P(AB) = P(A)P(B)$ ”.